The Value of A BANGLADESHI WOMAN'S TIME: AN ECONOMETRIC ANALYSIS

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## INTRODUCTION

## "management and care of the household".

## The System of National Accounts (SNA) is the internationally agreed standard set of recommendations on how to compile measures of economic activity.

## Accounts <br> System of National 2008



The System of National Accounts includes


Production \& trade of weapons


Prostitution


Women's unpaid labour

## Cake Model of the Economy

Monetised top parts of the cake account for all officially measured statistics of economic output


Unmonetised bottom parts of the cake subsidise the top parts with free labour and resources

- If women are not visible as contributors to a nation's economy, then they will not be visible in the distribution of benefits
- Value of housewife's labour can be used during divorce settlement or for estimating the extent of economic loss due to the wrongful death of a housewife
- Can complement time-use surveys to estimate economic contribution of housewives


## Failure to

 recognize the value ofwomen's unpaid work is failure to recognize the value of women
themselves

## Cosixililibili



END ALL VIOLENCE AGAINST AND EXPLOITATION OF WOMEN AND GIRLS

5 GENDER


VALUE UNPAID CARE
AND PROMOTE SHARED DOMESTIC
RESPONSIBILITIES

## Labour Force Participation Rate in Bangladesh

(ILO modelled estimates and projections)



Average number of hours spent per week doing household tasks

| MALE |  |
| :--- | :--- |
| Barisal | $8.2 \mathrm{hrs} /$ week |
| Chittagong | $9.4 \mathrm{hrs} /$ week |
| Dhaka | $9.4 \mathrm{hrs} /$ week |
| Khulna | $8.4 \mathrm{hrs} /$ week |
| Rajshahi | $6.9 \mathrm{hrs} /$ week |
| Rangpur | $7.6 \mathrm{hrs} /$ week |
| Sylhet | $12 \mathrm{hrs} /$ week |


| FEMALE |  |
| :--- | :--- |
| Barisal | $24 \mathrm{hrs} /$ week |
| Chittagong | $27 \mathrm{hrs} /$ week |
| Dhaka | $28 \mathrm{hrs} /$ week |
| Khulna | $27 \mathrm{hrs} /$ week |
| Rajshahi | $25 \mathrm{hrs} /$ week |
| Rangpur | $27 \mathrm{hrs} /$ week |
| Sylhet | $31 \mathrm{hrs} / w e e k$ |

On average, women spend 3 times more hours doing household work than men


Unpaid work refers to the production of goods or services that are consumed by those within or outside a household, but not for sale in the market


The reservation wage is the wage below which a person will not work, and in the labour/leisure context represents the value placed on an hour of lost leisure time

The imputed price or value of a good or service where such a price or value cannot be
accurately determined because of the absence of an ordinary market determined price

## THEORETICAL FRAMEWORK





## Utility $=f($ Leisure, Income $)$

# Time Allocation 

Time


## Employment Options



# LITERATURE REVIEW 



## Flawed assumptions of replacement cost method



## Review of past studies

| Authors | Year | Sampling strategy | Methodology | Results |
| :---: | :---: | :---: | :---: | :---: |
| Shamim Hamid | 1996 | Time budget survey through cluster sampling of 30 villages | Replacement cost | Inclusion of non-market work in official calculations would have increased national GDP 1989-90 by 29 per cent. |
| Debra Efroymson, Buddhadeb Biswas, and Shakila Ruma | 2007 | Survey of 630 men and women | i) Replacement cost <br> ii)Government salaries | i) US $\$ 131$ billion <br> ii) US \$ 152 billion |
| Debra Efroymson, Julia Ahmed, Shakila Ruma | 2013 | Survey of 630 men and women | i) Replacement cost <br> ii)Government salaries | i) US $\$ 227.93$ billion per year <br> ii) US $\$ 258.82$ billion per year |
| Rashed AI Mahmud <br> Titumir, K.M. <br> Mustafiqur Rahman | 2014 | Survey of 520 households in 7 districts | i) Replacement cost <br> ii) Opportunity cost | i) 3.25 percent of FY2012-2013 GDP <br> ii) 10.75 percent of FY20122013 GDP |
| Fahmida Khatun, Towfiqul Islam Khan, Shahida Pervin, Hosna Jahan | 2015 | Survey of 5,670 households | i) Replacement cost <br> ii) Willingness to accept | i) 76.8 percent of FY2013-2014 GDP <br> ii) 87.2 percent of FY2013-2014 GDP |

DATA

| Nationally | All variables <br> from QLFS 2015, <br> except inflation <br> (which is from |
| :--- | :--- |
| crosesentative sectional |  |
| data | Bangladesh <br> Bank) |

## Sampling strategy



Systematic random sampling of clusters of 24 households from each of the 1284 PSUs/EAs. Exactly 30,816 households are selected at this stage. Dataset has 503,756 observations

Random selection of 1284 PSUs/EAs from all of the 64 districts and 21 regional strata. Approximately 300,000 households are selected at this stage.

7 divisions: Barisal, Chittagong, Dhaka, Khulna, Rajshahi, Rangpur, and Sylhet. 3 kinds of localities: City Corporation, Urban, and Rural.

1284 primary sampling units (primary sampling units) or enumeration areas (EAs). Each PSU/EA had approximately 225 households.

## Variables used in the study

| Variable | Name | Definition | Mean | Std. Dev. |
| :---: | :---: | :---: | :---: | :---: |
| fwage | WAGE | weekly wage of married spouse present women in cash \& kind from both primary \& secondary job | 2479.155 | 1936.994 |
| feducation | EDUCATION | years of schooling for females up to 12 years | 4.259556 | 3.88348 |
|  |  | 0 = no education $6=$ class -VI |  |  |
|  |  | 1 = class-I 7 = class-VII |  |  |
|  |  | 2 = class-II 8 = class-VIII |  |  |
|  |  | 3 = class-III 9 = class-IX |  |  |
|  |  | 4 = class-IV $10=$ SSC |  |  |
|  |  | 5 = class-V $11=$ HSC |  |  |
| fexperience | EXPERIENCE | potential experience of females <br> experience $=$ [female age] - [6] (based on similar approach in Oaxaca, 1973) | 31.26629 | 11.93601 |
| fhours1 | LABOR SUPPLY | total number of hours worked per week at both primary \& secondary job | 31.88131 | 21.37751 |
| mwage | HUSBAND WAGE | weekly husband wage in cash \& kind from both primary \& secondary job | 2747.009 | 2154.392 |
| fCPI | GOODS PRICES | Consumer price index (CPI) | 219.9138 | . 1692468 |
| (BB Data) |  | $\mathrm{fCPI}=220.12$ if rural |  |  |
|  |  | fCPI $=219.37$ if urban |  |  |
|  |  | fCPI $=219.86$ if neither rural or urban (national) |  |  |
| fchildren | CHILDREN | number of children aged less than 6 per household | . 1664183 | . 4758965 |
| fasset | ASSETS | female asset dummy measured as total amount of land owned by households, measured in acres |  |  |
|  |  | fasset1; $1=$ no land, $0=$ all else | . 1890322 | . 3915362 |
|  |  | fasset2; $1=0.01-0.04$ acres land, $0=$ all else | . 4180809 | . 4932461 |
|  |  | fasset3; $1=0.05-2.49$ acres land, $0=$ all else | . 3434959 | . 4748777 |
|  |  | fasset4; $1=2.50-7.49$ acres land, $0=$ all else | . 0407771 | . 1977744 |
|  |  | fasset5; 1 = 7.50 acres or more land, $0=$ all else | . 0086139 | . 0924108 |



## Asking wage function

## $W^{*}=g\left(h, W_{m}, P, A, Z\right)(1)$

where,
$\mathrm{W}^{*}=$ asking wage rate (shadow price of time)
$\mathrm{h}=$ hours of work
$\mathrm{W}_{\mathrm{m}}=$ wage of husband
$\mathrm{P}=$ vector of goods prices
A = asset income of the household
Z $=$ number of children aged less than six

## Asking wage model specification

```
ln(\mp@subsup{W}{i}{*})=\mp@subsup{\beta}{0}{}+\mp@subsup{\beta}{1}{}\mp@subsup{h}{i}{}+\mp@subsup{\beta}{2}{}(\mp@subsup{W}{m}{})i+\mp@subsup{\beta}{3}{}\mp@subsup{P}{i}{}+\mp@subsup{\beta}{4}{}\mp@subsup{A}{i}{}+\mp@subsup{\beta}{5}{}\mp@subsup{Z}{i}{}+\mp@subsup{\varepsilon}{i}{}(3)
```

    where,
    \(\ln \left(W_{i}^{*}\right)=\) natural logarithm of asking wage rate
    (shadow price of time)
    \(\mathrm{h}=\) hours of work
    $\mathrm{W}_{\mathrm{m}} \quad=\quad$ wage of husband
$\mathrm{P} \quad=\quad$ vector of goods prices
A = asset income of the household
Z $\quad=\quad$ number of children aged less than six

## Market wage function

## $W=B(E, S)(2)$

where,
$\mathrm{W}=$ market wage rate (offered wage rate)
$\mathrm{E}=$ extent of labour market experience
$\mathrm{S}=$ number of years of schooling

Note that $B_{E}>0$ and $B_{S}>0$ from previous research

## Market wage model specification

## $\ln \left(W_{i}\right)=b_{0}+b_{1} S_{i}+b_{2} E_{i}+u_{i}(4)$

where,
$\ln \left(\mathrm{W}_{\mathrm{i}}\right)=$ natural logarithm of market wage rate (offered wage rate)
E $\quad=\quad$ extent of labour market experience
$\mathrm{S}=\quad$ number of years of schooling

## METHODOLOGY

> Equation (4) suffers from unobserved heterogeneity or the problem of omitted variables.

The effect of these unobserved variables is captured through the error terms, and so the errors of the equation (4) are correlated with the independent variables.

The underlying reason behind this is the fact that the samples used for estimating these equations were not randomly collected.


Market wages are only observed for women who are working.


Education

## Strict exogeneity assumption of the OLS model is

$$
E\left(\varepsilon_{i} \mid X\right)=0, \quad \forall i=1, \ldots, n
$$

Violation of the strict exogeneity assumption has several implications:

- $E\left(\varepsilon_{i}\right) \neq 0, \forall i=1, \ldots, n$
- (The unconditional mean of the error term ( $\varepsilon$ ) is not zero.)
- $E\left(X_{j k}, \varepsilon_{i}\right) \neq 0, \forall i j k=1, \ldots, n$
- (The independent variables (X) are not orthogonal to the errors ( $\varepsilon$ ) for all observations)
- $\operatorname{Cov}\left(X_{j k}, \varepsilon_{i}\right) \neq 0, \forall \boldsymbol{i j k}=1, \ldots, n$
- (The independent variables (X) and errors ( $\varepsilon$ ) are not uncorrelated for all observations.)

If the unobserved heterogeneity can be modelled separately, and the resulting information can be incorporated into the main model, then the problem can be resolved.

Heckman proposed that the specification of the original biased model could be improved by using the estimated values of the omitted variables as additional regressors.


Heckman outlined an ingenious two step estimation technique to correct sample selection bias (Heckman, 1979).

By doing so, the model could be estimated using ordinary least squares, without violating the strict exogeneity assumption.

In the first step, we model the factors that influence a woman's decision to work by using a probit model. The general form of the sample likelihood function for probit analysis is:

$$
\mathcal{L}=\prod_{i=1}^{T}\left[F\left(\phi_{i}\right)\right]^{1-d_{i}}\left[1-F\left(\phi_{i}\right)\right] d_{i}
$$

where, $d$ is a random variable, which is equal to one if the dependent variable is observed and equal to zero if the dependent variable is not observed.

Suppose we use a sample of $T$ married spouse present women, K of who work and T-K who do not work.
Then, in the case of our model, the aforementioned likelihood function becomes:

$$
\mathcal{L}=\prod_{i=1}^{K} j\left(h_{i}, \ln \left(W_{i}\right) \mid\left(W_{i}>W_{i}^{*}\right)_{h=0}\right) \cdot \operatorname{pr}\left(\left[W_{i}>W_{i}^{*}\right]_{h=0}\right) \times \prod_{i=K+1}^{T} \operatorname{pr}\left(\left[W_{i}<W_{i}^{*}\right]_{h=0}\right)
$$

# Inverse Mills Ratio $=$ standard normal probability distribution function standard normal cumulative distribution function 

$$
\lambda_{i}=\frac{f\left(\phi_{i}\right)}{1-F\left(\phi_{i}\right)}
$$

where,
$\boldsymbol{\lambda}=$ inverse Mills ratio
$\mathrm{f}=$ standard normal probability distribution function of the selection equation
$\mathrm{F}=$ standard normal cumulative distribution function of the selection equation.

For our model, the Inverse Mills Ratio can be defined as:

$$
\lambda=j\left(h_{i}, \ln \left(W_{i}\right) \left\lvert\,\left(W_{i}^{*}<W_{i}\right)_{h=0}=\frac{n\left(h_{i}, \ln \left(W_{i}\right)\right)}{\operatorname{pr}\left(\left[W_{i}>W_{i}^{*}\right]_{h=0}\right)} \because \varepsilon_{i}\right., u_{i} \sim N(0)\right.
$$

By using this Inverse Mills Ratio in our original likelihood function, we can further simplify it to:

$$
\mathcal{L}=\prod_{i=1}^{K} n\left(h_{i}, \ln \left(W_{i}\right)\right) \prod_{i=K+1}^{T} \operatorname{pr}\left(\left[W_{i}<W_{i}^{*}\right]_{h=0}\right)
$$

We now maximize this likelihood function with respect to the parameters of the model, including the variances and covariances of the errors in equations (3) and (4) to get consistent, asymptotically unbiased, and efficient parameter estimates which are asymptotically normally distributed.

Thus, our selection bias corrected now becomes:

$$
\ln \left(W_{i}\right)=b_{0}+b_{1} S_{i}+b_{2} E_{i}+b_{3} \lambda_{i}+u_{i}(5)
$$

## RESULTS

## Results from Ordinary Least Squares Estimation

VARIABLES
Natural log of female wage

| Education | $0.0458^{* * *}$ |
| :--- | :---: |
|  | $(0.00134)$ |
| Experience | $0.00303^{* * *}$ |
|  | $(0.000513)$ |
| Constant | $7.426^{* *}$ |
|  | $(0.0175)$ |
| Prob > F | 0.0000 |
| R-squared | 0.129 |
| Adj R-squared | 0.1289 |

Note: (i) Standard errors in parentheses (ii) *** $p<0.01$, ** $p<0.05$, * $p<0.1$

## Ramsey Regression Specification Error Test (RESET)

Null hypothesis = model is correctly specified Alternative hypothesis = model is incorrectly specified

Decision rule: if $p<0.05$ then the model is incorrectly specified

$$
\begin{aligned}
& \text { Results from Ramsey RESET test } \\
& F(3,7949)=120.01 \\
& \text { Prob }>F=0.0000
\end{aligned}
$$

Interpretation: The model is incorrectly specified

## Link Test

- Link Test is based on the idea that if a regression is properly specified, one should not be able to find any additional independent variables that are significant except by chance.
- Link Test creates two new variables, the variable of prediction, and the variable of squared prediction.
- We wouldn't expect the squared prediction to be a significant predictor if our model is specified correctly.

| Results from Link Test |  |
| :--- | :---: |
| VARIABLES | Infwage |
| Prediction | $-47.66363^{* * *}$ |
|  | $(3.02233)$ |
| Squared | $3.146831^{* * *}$ |
| prediction |  |
|  | $(0.1954301)$ |
| Constant | $188.0395^{* * *}$ |
|  | $(11.68005)$ |
| Prob > F | 0.0000 |
| R-squared | 0.1566 |
| Adj R-squared | 0.1564 |

## Interpretation: The model is incorrectly specified

## Variance Inflation Factor

Variance inflation factor measures the linear association between an independent variable and all other independent variables.
Decision rule:
VIF > 10 : perfect multicollinearity is highly likely
$5<$ VIF < 10 : perfect multicollinearity is somewhat likely $0<\mathrm{VIF}<5$ : perfect multicollinearity is unlikely


Interpretation: Perfect multicollinearity is unlikely

## Breusch-Pagan and Cook-Weisberg Test

Null hypothesis = homoskedastic,
Alternative hypothesis = heteroskedaticity
Decision rule: if $p<0.05$ then there is heteroskedasticity.

## Results from Breusch-Pagan / Cook-Weisberg Test for Heteroskedasticity

| $\operatorname{chi2}(1)=0.00$ | $F(1,7953)=0.00$ |
| :--- | :--- |
| Prob $>$ chi2 $=0.9930$ | Prob $>F=0.9970$ |

Interpretation: There is no heteroskedasticity

## Graphical Check of Heteroskedasticity

Scatter Plot of Residuals vs Fitted Values


## White Test

Breusch-Pagan (1979) and Cook-Weisberg (1983) test for heteroskedasticity assumes that the heteroskedasticity is a linear function of the independent variables.

The White test allows the heteroskedasticity process to be a function of one or more independent variables. It allows the independent variable to have a non-linear and interactive effect on the error variance.

Null hypothesis = homoscedastic; Alternative hypothesis = heteroskedaticity Decision rule: if $p<0.05$ then there is heteroskedasticity.

## Results from White's Test for Heteroskedasticity <br> chi2(5) $=39.34$ <br> Prob $>$ chi2 $=0.0000$

Interpretation: There is heteroskedasticity

## Graphical Check of Normality of Errors (PP \& QQ)



PP plot is more sensitive in the center; QQ plot is more sensitive at the two tails

## Graphical Check of Normality of Errors (Histogram and Kernel Density)




## Shapiro Wilk Test

Null hpyothesis = errors normal
Alternative hypothesis = errors not normal
Decision rule: If $p$ value $<0.05$ then reject null hypothesis that errors are normal. If $p$ value $>0.05$ then cannot reject null hypothesis that errors are normal.

## Results from Shapiro-Wilk Test for Normal Data

| Variable | W | V | z | Prob>z |
| :--- | :---: | :--- | :---: | :--- |
| Residual | 0.90892 | 371.790 | 15.744 | 0.00000 |

Interpretation: The joint distribution of the errors is not normal.

## Graphical Check of Outliers in Regression Model




## Summary of Post-estimation Diagnostic Tests of OLS Estimation

| Test | Description | Result |
| :--- | :--- | :--- |
| Ramsey Regression <br> Specification Error Test | Test of model specification | Model is incorrectly <br> specified; there is at least <br> one omitted variable |
| Link Test | Test of model specification | Model is incorrectly <br> specified |
| Variance Inflation Factor | Test of multicollinearity | Perfect multicollinearity is <br> unlikely |
| Breusch-Pagan and Cook- <br> Weisberg Test | Test of heteroskedasticity | There is no <br> heteroskedasticity |
| White Test | Test of heteroskedasticity | There is heteroskedasticity <br> Shapiro Wilk Test |
|  | Test of normality of errors | The errors are not normally <br> distributed |

## Results from Heckman Two-step Estimation

|  | Probit | Heckman |
| :---: | :---: | :---: |
| VARIABLES | Natural log of female wage | Natural log of female wage |
| Husband wage | $\begin{aligned} & -3.67 * 10^{-06} \\ & \left(5.43^{\star 1} 0^{-06}\right) \end{aligned}$ |  |
| Hours | $\begin{gathered} 0.0342^{* * *} \\ (0.000667) \end{gathered}$ |  |
| Goods prices | $\begin{aligned} & 0.700^{\star * *} \\ & (0.0810) \end{aligned}$ |  |
| NO land | $0.616^{* *}$ |  |
| 0.01-0.04 acres land | $\begin{gathered} 0.260 \\ (0.311) \end{gathered}$ |  |
| 0.05-2.49 acres land | $\begin{array}{r} -0.133 \\ (0.311) \end{array}$ |  |
| 2.50-7.49 acres land | $\begin{array}{r} -0.259 \\ (0.335) \end{array}$ |  |
| 7.50 acres or more land | (omitted) |  |
| Children | $\begin{gathered} -0.0793^{* * *} \\ (0.0299) \end{gathered}$ |  |
| Education |  | $\begin{aligned} & 0.0348^{\star * *} \\ & (0.00203) \end{aligned}$ |
| Experience |  | $\begin{gathered} 0.00108 \\ (0.000810) \end{gathered}$ |
| Lambda |  | $\begin{gathered} -0.0897 * * * \\ (0.0172) \end{gathered}$ |
| Rho |  | -0.19942 |
| sigma |  | 0.44970462 |
| Constant | $\begin{gathered} -155.8^{* * *} \\ (17.83) \end{gathered}$ | $\begin{aligned} & 7.592^{\star * *} \\ & (0.0272) \end{aligned}$ |
| LR chiz (8) | 3803.82 |  |
| Prob $>$ chiz | 0.0000 |  |
| Wald chiz (2) |  | 324.93 |
| Prob $>$ chiz |  | 0.0000 |

## Estimated Reservation Wage of the Average Bangladeshi Housewife

| Years of | Average Monthly <br> Reservation Wage of <br> Housewife (in BDT) | Reservation Wage of <br> Housewife as \% of <br> Husband's Average <br> Monthly Wage | Standard error | P>\|z| |
| :---: | :---: | :---: | :---: | :---: |
| 0 | 8502.88 | $72.22 \%$ | 0.026569 | 0.000 |
| 1 | 8803.63 | $74.78 \%$ | 0.0254757 | 0.000 |
| 2 | 9115.00 | $77.42 \%$ | 0.0245018 | 0.000 |
| 3 | 9437.40 | $80.16 \%$ | 0.023662 | 0.000 |
| 4 | 9771.20 | $82.99 \%$ | 0.0229711 | 0.000 |
| 5 | 10116.81 | $85.93 \%$ | 0.0224428 | 0.000 |
| 6 | 10474.63 | $88.97 \%$ | 0.0220889 | 0.000 |
| 7 | 10845.12 | $92.12 \%$ | 0.0219177 | 0.000 |
| 8 | 11228.70 | $95.38 \%$ | 0.0219335 | 0.000 |
| 9 | 11625.86 | $98.75 \%$ | 0.022136 | 0.000 |
| 10 | 12037.06 | $102.24 \%$ | 0.02252 | 0.000 |
| 11 | 12462.81 | $105.86 \%$ | 0.0230766 | 0.000 |

## Conclusion

Essential for the economy

Priceless; not worthless

What you can measure, you treasure

Recognition is a stepping stone

## Women's

 unpaid labour
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## Disclaimer

This research was submitted in partial fulfillment of the requirements of the degree of Master of Science in Economics at North South University (NSU), Bangladesh. The views expressed in this presentation are the author's own, and do not represent those of North South University (NSU), The Centre for Policy Dialogue (CPD), or of any other organisation or individual. The author assumes full responsibility for any unintentional errors or omissions contained in this presentation and its supporting research paper.

## "I neither know nor think that I know"

- Socrates


## THANK YOU



